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# Exponential Growth State Space model version date 120815.R:
#
# Program for calculating maximum likelihood (ML) or restricted maximum
# likelihood (REML) estimates of unknown parameters for the Exponential
# Growth State Space (EGSS) model of stochastic population growth.
# The model is
#
#  $X(t) = \mu + E(t)$ 
#           with  $E(t) \sim \text{normal}(0, \text{ssq})$ ,
#  $Y(t) = X(t) + F(t)$ 
#           with  $F(t) \sim \text{normal}(0, \text{tsq})$ .
#
# Here  $X(t)$  is log-population abundance,  $Y(t)$  is observed or estimated value
# of  $X(t)$ ,  $x_0$ ,  $\mu$ ,  $\text{ssq}$ ,  $\text{tsq}$  are parameters. The parameter  $\text{ssq}$  is the
# variance of the process noise, and  $\text{tsq}$  is the variance of the observation
# error.
# The model takes population abundance  $N(t) = \exp(X(t))$  to be governed by a
# stochastic, density independent model, with the observed
# abundances  $O(t) = N(t) \cdot \exp(F(t))$  arising from lognormal sampling error.
#
# User provides time series of observed population abundances  $o(0)$ ,  $o(1)$ ,
# ...,  $o(q)$ , which are log-transformed by the program into  $y(0)$ ,  $y(1)$ , ...,
#  $y(q)$ , assumed to be a time series realization of  $Y(t)$ . Likelihood
# function of  $y(0)$ ,  $y(1)$ , ...,  $y(q)$  is that of a multivariate normal
# distribution. The observation times  $t_0$ ,  $t_1$ ,  $t_2$ , ...,  $t_q$  can have
# unequal intervals.
#
# Program computes initial parameter values for iterations. The program
# should be re-run for several sets of initial values, as the likelihood
# function for the model frequently has multiple local maxima.
#
# An alternative program, in SAS, is available as an online appendix to
# Staples et al. (2004).
#
# This program written by Brian Dennis (Dept Fish and Wildlife Sciences,
# Univ Idaho, Moscow, Idaho, 83844-1136 USA).
#
# Citations:
#   Staples et al. 2004. Ecology.
#   Dennis et al. 2006. Ecological Monographs.
#   Humbert et al. 2009. Oikos.
#
#-----
#           USER INPUT SECTION
#-----
# User supplies time series data here. User can substitute R statements to
# read population abundance data from a file into the vector "Observed.t".
# Times of observation are entered into the vector "Time.t".
# Observed.t=c(18,10,9,14,17,14,5,10,9,5,11,11,4,5,4,8,2,3,9,2,4,7,4,1,2,
# 4,11,11,9,6) # No zeros! (With zeros, you must use another model)
# Time.t=c(0,1,2,3,4,5,6,7,8,9,10,11,12,13,14,15,16,17,18,19,20,21,22,23,24,
# 25,26,27,28,29) # Initial time can be nonzero.
#
# Example data are American Redstart counts from North American Breeding
# Bird Survey, record # 02014 3328 08636, 1966-95 (Table 1 in Dennis
# et al. 2006).

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# Polar bear, from Canada
Observed.t=c(1213,1187,1295,1067,1038,969,1233,1027,1236,1039,1055,
  970,1145,875,1144,968,965,869) # No zeros!
Time.t=c(1987,1988,1989,1990,1991,1992,1993,1994,1995,1996,1997,1998,
  1999,2000,2001,2002,2003,2004)

#-----
#          PROGRAM INITIALIZATION SECTION
#-----
library(MASS) # loads miscellaneous functions (ginv, etc.)
T.t=Time.t-Time.t[1] # Time starts at zero.
Y.t=log(Observed.t) # Log-transform the observations.
q=length(Y.t)-1 # Number of time series transitions, q.
qp1=q+1 # q+1 gets used a lot, too.
S.t=T.t[2:qp1]-T.t[1:q] # Time intervals.
m=rep(1,qp1) # Will contain Kalman means for Kalman calculations.
v=rep(1,qp1) # Will contain variances for Kalman calculations.

#-----
#          SECTION FOR DEFINING ML & REML LOG-LIKELIHOODS
#-----

# ML objective function "negloglike.ml" is negative of log-likelihood;
# the Nelder-Mead optimization routine in R, "optim", is a minimization
# routine. The ML objective function uses equations 24-26 from Dennis et
# al. (2006). The three function arguments are: theta, vector of
# parameters
# (transformed to the real line), yt, vector of time series observations,
# and
# tt, vector of observation times.
negloglike.ml=function(theta,yt,tt) {
  muu=theta[1]
  sigmasq=exp(theta[2]) # Constrains ssq > 0.
  tausq=exp(theta[3]) # Constrains tsq > 0.
  xzero=theta[4]
  q=length(yt)-1
  qp1=q+1
  vx=matrix(0,qp1,qp1)
  for (ti in 1:q) {
    vx[(ti+1),(ti+1):qp1]=matrix(1,1,(qp1-ti))*tt[ti+1]
    vx[(ti+1):qp1,(ti+1)]=t(vx[(ti+1),(ti+1):qp1])
  }
  Sigma.mat=sigmasq*vx
  Itausq=matrix(rep(0,(qp1*qp1)),nrow=qp1,ncol=qp1)
  diag(Itausq)=rep(tausq,qp1)
  V=Sigma.mat+Itausq
  mu=xzero+muu*tt
  ofn=((qp1)/2)*log(2*pi)+(0.5*log(det(V)))+
    (0.5*((yt-mu)%*%ginv(V)%*%(yt-mu)))
  return(ofn)
}

# REML objective function "negloglike.reml" is negative of log-likelihood
# for second differences of the log-scale observations. The REML objective
# function uses equations A22-A25 from Humbert et al. (2009). The three
# function arguments are: theta, vector of parameters (transformed to the
# real line), yt, vector of time series observations (log scale), and

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# tt, vector of observation times. Function performs the differencing.
negloglike.reml=function(theta,yt,tt) {
  sigsq=exp(theta[1])          # Constrains ssq > 0.
  tausq=exp(theta[2])         # Constrains tsq > 0.
  q=length(yt)-1
  qp1=q+1
  vx=matrix(0,qp1,qp1)
  for (ti in 1:q) {
    vx[(ti+1),(ti+1):qp1]=matrix(1,1,(qp1-ti))*tt[ti+1]
    vx[(ti+1):qp1,(ti+1)]=t(vx[(ti+1),(ti+1):qp1])
  }
  Sigma.mat=sigsq*vx
  Itausq=matrix(rep(0,(qp1*qp1)),nrow=qp1,ncol=qp1)
  diag(Itausq)=rep(tausq,qp1)
  V=Sigma.mat+Itausq
  ss=tt[2:qp1]-tt[1:q]
  D1mat=cbind(-diag(1/ss),matrix(0,q,1))+cbind(matrix(0,q,1),diag(1/ss))
  D2mat=cbind(-diag(1,q-1),matrix(0,q-1,1))+
    cbind(matrix(0,q-1,1),diag(1,q-1))
  Phi.mat=D2mat%*%D1mat%*%V%*%t(D1mat)%*%t(D2mat)
  wt=(yt[2:qp1]-yt[1:q])/ss
  ut=wt[2:q]-wt[1:q-1]
  ofn=(q/2)*log(2*pi)+(0.5*log(det(Phi.mat)))+
    (0.5*(ut%*%ginv(Phi.mat)%*%ut))
  return(ofn)
}

#-----
#          SECTION FOR CALCULATING EGOE AND EGPN ESTIMATES
#          (FOR USE AS INITIAL VALUES)
#-----
Ybar=mean(Y.t)
Tbar=mean(T.t)
mu.egoe=sum((T.t-Tbar)*(Y.t-Ybar))/sum((T.t-Tbar)*(T.t-Tbar))
x0.egoe=Ybar-mu.egoe*Tbar
ssq.egoe=0
Yhat.egoe=x0.egoe+mu.egoe*T.t
tsq.egoe=sum((Y.t-Yhat.egoe)*(Y.t-Yhat.egoe))/(q-1)

Ttr=sqrt(S.t)
Ytr=(Y.t[2:qp1]-Y.t[1:q])/Ttr
mu.egpn=sum(Ttr*Ytr)/sum(Ttr*Ttr)
Ytrhat=mu.egpn*Ttr
ssq.egpn=sum((Ytr-Ytrhat)*(Ytr-Ytrhat))/(q-1)
tsq.egpn=0
x0.egpn=Y.t[1]

mu0=(mu.egoe+mu.egpn)/2      # Initial values for EGSS are averages
ssq0=ssq.egpn/2             #   of EGOE and EGPN values
tsq0=tsq.egoe/2            #   --
x00=x0.egoe                 #   --

#-----
#          SECTION FOR CALCULATING ML & REML PARAMETER ESTIMATES
#-----

# The ML estimates.

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EGSSml=optim(par=c(mu0,log(ssq0),log(tsq0),x00),
  negloglike.ml, NULL, method="Nelder-Mead", yt=Y.t, tt=T.t)
params.ml=c(EGSSml$par[1], exp(EGSSml$par[2]), exp(EGSSml$par[3]),
  EGSSml$par[4])
lnlike.ml=-EGSSml$value[1]
AIC.egss=-2*lnlike.ml+2*length(params.ml)

mu.ml=params.ml[1]          # These are the ML estimates.
ssq.ml=params.ml[2]        #          --
tsq.ml=params.ml[3]        #          --
x0.ml=params.ml[4]         #          --

# The REML estimates.
EGSSreml=optim(par=c(log(ssq0),log(tsq0)),
  negloglike.reml, NULL, method="Nelder-Mead", yt=Y.t, tt=T.t)
params.reml=c(exp(EGSSreml$par[1]), exp(EGSSreml$par[2]))
ssq.reml=params.reml[1]    # These are the REML estimates.
tsq.reml=params.reml[2]    #          --

vx=matrix(0, qp1, qp1)
for (ti in 1:q) {
  vx[(ti+1), (ti+1):qp1]=matrix(1, 1, (qp1-ti))*T.t[ti+1]
  vx[(ti+1):qp1, (ti+1)]=t(vx[(ti+1), (ti+1):qp1])
}

Sigma.mat=ssq.reml*vx
Itausq=matrix(rep(0, (qp1*qp1)), nrow=qp1, ncol=qp1)
diag(Itausq)=rep(tsq.reml, qp1)
V=Sigma.mat+Itausq
Dlmat=cbind(-diag(1/S.t), matrix(0, q, 1))+cbind(matrix(0, q, 1), diag(1/S.t))
Vlmat=Dlmat%*%V%*%t(Dlmat)
W.t=(Y.t[2:qp1]-Y.t[1:q])/S.t
j1=matrix(1, q, 1)
Vlinv=ginv(Vlmat)
mu.reml=(t(j1)%*%Vlinv%*%W.t)/(t(j1)%*%Vlinv%*%j1)
j=matrix(1, qp1, 1)
Vinv=ginv(V)
x0.reml=(t(j)%*%Vinv%*%(Y.t-mu.reml*T.t))/(t(j)%*%Vinv%*%j)
Var_mu.reml=1/(t(j1)%*%Vlinv%*%j1)          # Variance of mu
mu_hi.reml=mu.reml+1.96*sqrt(Var_mu.reml)   # 95% CI for mu
mu_lo.reml=mu.reml-1.96*sqrt(Var_mu.reml)   #          --

# Calculate estimated population sizes for EGSS model
# with Kalman filter, for plotting.
#
# Choose ML or REML estimates here (by commenting out the unwanted).
# mu=mu.ml;  ssq=ssq.ml;  tsq=tsq.ml;  x0=x0.ml;
mu=mu.reml;  ssq=ssq.reml;  tsq=tsq.reml;  x0=x0.reml;

m[1]=x0          # Initial mean of Y(t).
v[1]=tsq         # Initial variance of Y(t).

for (ti in 1:q) {
  # Loop to generate estimated population abundances
  # using Kalman filter (see equations 6 & 7,

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# Dennis et al. (2006).
m[ti+1]=mu+(m[ti]+((v[ti]-tsq)/v[ti])*(Y.t[ti]-m[ti]))
v[ti+1]=tsq*((v[ti]-tsq)/v[ti])+ssq+tsq
}

# The following statement calculates  $\exp\{E[X(t) | Y(t), Y(t-1), \dots, Y(0)]\}$ ;
# see equation 54 in Dennis et al. (2006).
Predict.t=exp(m+((v-tsq)/v)*(Y.t-m))

# Plot the data & model-fitted values
plot(Time.t,Observed.t,xlab="time",ylab="population abundance",
      type="o",pch=1,cex=1.5) # Population data are circles.
par(lty="dashed") # Estimated abundances are dashed line.
points(Time.t,Predict.t, type="l", lwd=1)

# Print the parameter estimates
parms.egoe=c(mu.egoe,ssq.egoe,tsq.egoe,x0.egoe) # Collect for printing
parms.egpn=c(mu.egpn,ssq.egpn,tsq.egpn,x0.egpn) # --
parms.reml=c(mu.reml,ssq.reml,tsq.reml,x0.reml) # --
parms.ml=c(mu.ml,ssq.ml,tsq.ml,x0.ml) # --
names=c("mu","ssq","tsq","x0") # --
types=c("EGOE","EGPN","EGSS-ML","EGSS-REML") # --
matrix(cbind(parms.egoe,parms.egpn,parms.ml,parms.reml),
       nrow=4,ncol=4,byrow=TRUE,dimnames=list(types,names)) # Print stuff

matrix(cbind(mu_lo.reml,mu_hi.reml),nrow=1,ncol=2,byrow=TRUE,
       dimnames=list("95% CI for MU",c("LO","HI"))) # Print stuff

matrix(cbind(lnlike.ml,AIC.egss),nrow=1,ncol=2,byrow=TRUE,
       dimnames=list("EGSS ML RESULTS",c("LN-LIKELIHOOD","AIC"))) # Print stuff

```