

## Are Guinea Pigs Rodents? The Importance of Adequate Models in Molecular Phylogenetics

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The monophyly of Rodentia has repeatedly been challenged based on several studies of molecular sequence data. Most recently, D'Erchia *et al.* (1996) analyzed complete mtDNA sequences of 16 mammals and concluded that rodents are not monophyletic. We have reanalyzed these data using maximum-likelihood methods. We use two methods to test for significance of differences among alternative topologies and show that (1) models that incorporate variation in evolutionary rates across sites fit the data dramatically better than models used in the original analyses, (2) the mtDNA data fail to refute rodent monophyly, and (3) the original interpretation of strong support for nonmonophyly results from systematic error associated with an oversimplified model of sequence evolution. These analyses illustrate the importance of incorporating recent theoretical advances into molecular phylogenetic analyses, especially when results of these analyses conflict with classical hypotheses of relationships.

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**KEY WORDS:** inconsistency; maximum likelihood; molecular systematics; rodents; rate heterogeneity.

### INTRODUCTION

The assertions made in several molecular phylogenetic studies (Graur *et al.*, 1991; Li *et al.*, 1992; Ma *et al.*, 1993) have led to the growing acceptance of the conclusion that the order Rodentia is not monophyletic, in spite of the facts that these data sets essentially provide no significant refutation of the classical hypothesis (e.g., Hasegawa *et al.*, 1992; Cao *et al.*, 1994), and other molecular studies actually support rodent monophyly (Martignetti and Brosius, 1993; Porter *et al.*, 1996). Recently, D'Erchia *et al.* (1996) suggested that their phylogenetic analyses of complete mtDNA sequences of 16 mammalian species firmly establish that the guinea pig is not a rodent, based on its placement as a sister taxon to a clade containing Lagomorpha, Carnivora, Primates, Perissodactyla, and Artiodactyla (including cetaceans), rather than in a clade with mouse and rat. They claim that this placement both is consistent across phylogenetic reconstruction methodologies and is supported by "very significant" bootstrap values. Because nonmonophyly of the rodents would imply a remarkable amount of convergence in mor-

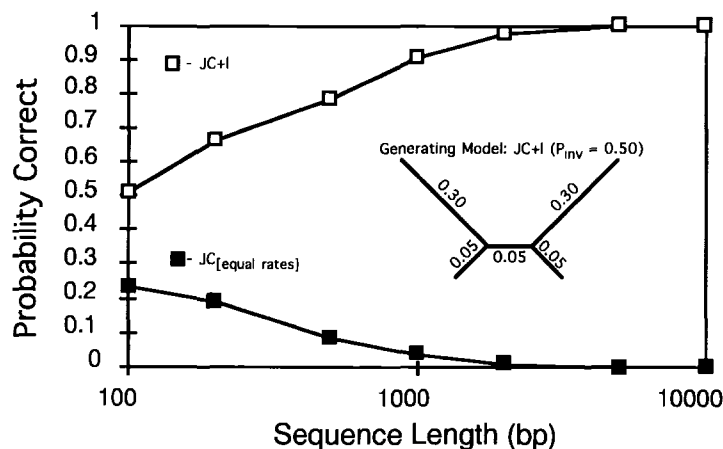
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phology [including the masticatory apparatus, cranial and postcranial skeletal characters, and placentation patterns; reviewed by Luckett and Hartenberger (1993)], the possibility that the results of D'Erchia *et al.* (1996) stem from systematic error in the mtDNA analyses must be explored.

The purpose of this paper is to assess the degree to which the conclusions of D'Erchia *et al.* (1996) are supported by their data. These authors estimated evolutionary trees using parsimony, distance, and maximum-likelihood methods. A frequent criticism of parsimony methods is their susceptibility to systematic error associated with "long-branch attraction" and related phenomena (e.g., Felsenstein, 1978; Hendy and Penny, 1989). Distance and likelihood methods have the advantage of being based on explicit models of evolutionary change, but recent studies have illustrated that even model-based methods are not immune to being inconsistent estimators of phylogeny when their assumptions are strongly violated (Gaut and Lewis, 1995; Waddell, 1995; Yang, 1996). The inconsistency of likelihood analysis under an oversimplified evolutionary model is demonstrated in Fig. 1. Sequences were simulated on the tree shown under a Jukes-Cantor model (Jukes and Cantor, 1969) with 50% invariable sites (JC+I). The upper curve represents the probability of inferring the true tree with increasing sequence length when the reconstruction model fits the data, and the lower curve represents the probability of inferring the true tree with increasing sequence length using likelihood under an assumption of an equal-rates Jukes-Cantor model. Clearly, maximum likelihood is inconsistent when all sites are incorrectly assumed to be free to vary.

The model-based analyses that D'Erchia *et al.* (1996) conducted on nucleotide sequences used an evolutionary model that allows both unequal base frequencies and different probabilities for each of the six possible transformations ( $A \rightleftharpoons C$ ,  $A \rightleftharpoons G$ ,



**Fig. 1.** Maximum-likelihood can be an inconsistent estimator of phylogeny. One thousand replicate data sets of several different sequence lengths were simulated on the tree shown under a JC+I model of sequence evolution with 50% invariable sites ( $P_{inv} = 0.5$ ). Each data set was then analyzed under both the appropriate JC+I model (open squares) and an incorrect equal-rates JC model (filled squares). Under the incorrect assumption of equal rates, the probability of inferring the correct tree is zero for long sequences (>2000 bp).

A ↔ T, C ↔ G, C ↔ T, G ↔ T). However their model assumes all sites are variable and evolve at the same rate. This assumption is clearly violated in mammalian mtDNA (e.g., Yang *et al.*, 1994; Sullivan *et al.*, 1995). Our reanalysis of their “protein supergene” data set (combined data from all mitochondrial protein genes) reveals that the evidence for excluding the guinea pig from the rodent clade is strongly overstated and is attributable to D’Erchia and co-workers’ use of an oversimplified model.

## METHODS

We chose to focus on the “protein supergene” data set because it provided stronger apparent refutation of rodent monophyly than the “rRNA supergene” data set. We omitted third-codon positions because these sites strongly violate the assumption of stationary base frequencies made by nearly all phylogenetic methods (either explicitly or implicitly); this data set is therefore very similar to the data set containing only nonsynonymous substitutions analyzed by D’Erchia *et al.* (1996). Because the MOLPHY program (Adachi and Hasegawa, 1996) used in D’Erchia and co-workers’ maximum-likelihood analysis does not allow for among-site rate heterogeneity, we performed maximum-likelihood analyses at the nucleotide (rather than amino acid) level. All analyses were conducted using test versions of the PAUP\* computer program (4.0d46–4.0d53) written by one of us (D.L.S.).

Our emphasis on maximum-likelihood methods is motivated by two considerations. First, it has been shown that maximum likelihood is a consistent estimator of phylogeny over a larger set of conditions than is parsimony (e.g., Huelsenbeck, 1995). Second, a major advantage of likelihood relative to parsimony or distance methods is that the likelihood score provides an objective criterion of goodness-of-fit between model and data that is comparable across models. This property provides a means for choosing an appropriate reconstruction model for phylogenetic analysis. Under parsimony, the optimality criterion (tree length) is not directly comparable across weighting schemes; this makes choice of weighting schemes (including the choice of equal weights) under the parsimony framework somewhat arbitrary.

We examined four substitution models: Jukes–Cantor [JC (Jukes and Cantor, 1969)], Kimura two-parameter [K2P (Kimura, 1980)], Hasegawa–Kishino–Yano [HKY-85 (Hasegawa *et al.*, 1985)], and general time-reversible [GTR; equals REV of Yang (1994a)]. The JC model assumes that all six transformations (A ↔ C, A ↔ G, A ↔ T, C ↔ G, C ↔ T, G ↔ T) have equal probability and that all four nucleotides are present in equal frequencies. The K2P model also assumes equal base frequencies but allows different probabilities for transitions and transversions (i.e., a transition bias). The HKY-85 model also allows for a transition bias and, further, relaxes the assumption of equal base frequencies. The GTR model allows unequal base frequencies and allows a unique probability for each of the six possible transformations.

In addition, four models of among-site rate heterogeneity were examined: (1) equal rates assumed at all sites; (2) a proportion of sites estimated to be invariable, with equal rates assumed at variable sites [“I” (Hasegawa *et al.*, 1985)]; (3) rates at all sites assumed to follow a discrete approximation of the gamma distribution [“Γ” (Yang, 1994b)]; and (4) some sites assumed to be invariable, with gamma-distributed rates at variable sites [“I + Γ” (Gu *et al.*, 1995; Waddell and Penny, 1996)]. Thus, 16 models

of sequence evolution were examined (four substitution models, each with four rate-heterogeneity models), each of which is a special case of the most parameter-rich model, GTR+I+ $\Gamma$ . The assumptions made by each of these models are compared in the Appendix [see Swofford *et al.* (1996) for a more detailed description of models].

Separate heuristic tree searches (stepwise addition of taxa, 10 random input orders, and TBR branch swapping) were conducted under the equal-rates GTR model [equivalent to the model used by D'Erchia *et al.* (1996)] and under the heterogeneous-rates model with the best fit to the data (GTR+I+ $\Gamma$ ; see below). An initial search was conducted with model parameters fixed to values estimated using the topology of D'Erchia *et al.* (1996). These parameters were then reoptimized on the resulting tree to refine the model further for subsequent tree searches. Both unconstrained searches and searches constrained for rodent monophyly were conducted, and the significance of differences in likelihood scores of alternative topologies was examined using the test of Kishino and Hasegawa (1989) and a simulation method similar to the parametric bootstrap (Huelsenbeck *et al.*, 1996).

The Kishino–Hasegawa test uses the standard error of differences in single-site likelihoods between two trees to estimate the significance of an observed difference between them [under the assumption that the distribution of single-site likelihood differences approximates a normal distribution (Kishino and Hasegawa, 1989)]. The simulation method estimates the probability of observing a particular result under a model of sequence evolution that is estimated from the data, assuming that some hypothesis is true. In this case, it is important to ask how often a tree on which rodents are monophyletic would be expected to generate data that appear to reject rodent monophyly. The best-fit tree (under the likelihood criterion) supporting rodent monophyly was used as the model (true) tree and 100 replicate data sets were simulated under the best-fit (GTR+I+ $\Gamma$ ; see below) model estimated from the original data. Each of these simulated data sets was then subjected to heuristic searches under parsimony (10 random input orders) and likelihood (simple addition sequence) using the GTR+I+ $\Gamma$  and GTR equal-rates model (see below). The proportion of replicates in which rodents are nonmonophyletic represents the probability of incorrectly inferring nonmonophyly if the best rodent monophyly tree were the true tree.

## RESULTS AND DISCUSSION

### Model Choice

The log-likelihood score of the tree of D'Erchia *et al.* (1996) under the GTR model (equal rates; equivalent to their model) is  $-49,028.14$ , whereas its score under the GTR+I+ $\Gamma$  is  $-45,593.28$  (see Appendix for likelihood of this tree under all models tested). This very large difference in likelihood score (3457.36 log-likelihood units) demonstrates that the fit between the data and the reconstruction model is dramatically improved by assuming that a proportion of sites are not free to vary and that rates at the remaining sites follow a gamma distribution (with shape parameter estimated from the data). The significance of this improvement can be evaluated using a likelihood-ratio test. The test statistic is twice the difference in log-likelihood and this can be compared

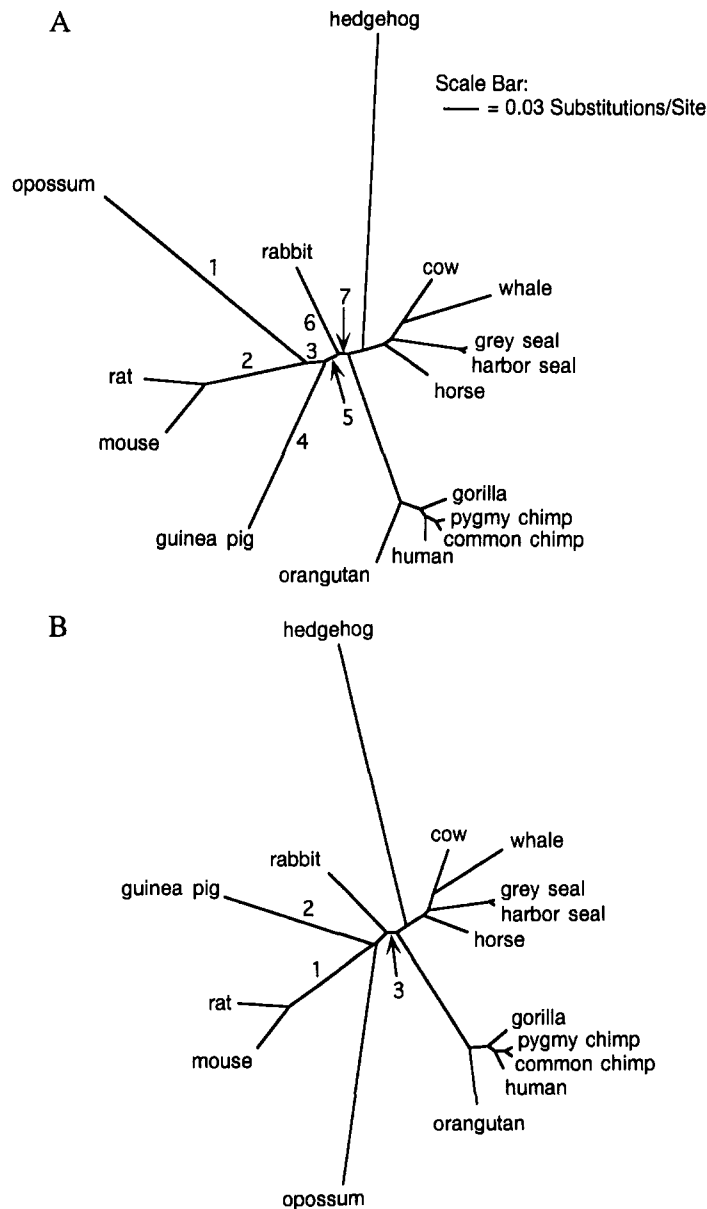
to the chi-squared distribution, with degrees of freedom equal to the difference in the number of free parameters between the two models (two in this case: the proportion of sites that are invariable and the gamma-distribution shape parameter). The value of the test statistic is 6914.72, whereas the critical value (at the 0.001 significance level) is 13.82. Although the assumptions of this test may not be strictly met (Goldman, 1993; but see Yang *et al.*, 1995), it nevertheless highlights the dramatic improvement in fit associated with allowing for heterogeneous rates.

### Phylogenetic Analyses

#### *Maximum-Likelihood Analyses Using Different Models Produce Different Trees*

Under the equal-rates GTR model, the tree shown by D'Erchia *et al.* (1996) is indeed the maximum-likelihood tree, but the maximum-likelihood tree under the more appropriate heterogeneous-rates GTR model (Fig. 2A) no longer supports the basal placement of the hedgehog. Furthermore, the hedgehog branch is remarkably long, longer even than the branch leading to the outgroup (opossum). The possibility of artifactual results due to long-branch attraction is well-known in parsimony analysis (Felsenstein, 1978), but the same problem can affect distance and maximum-likelihood analyses if the amount of change in long branches is systematically underestimated. This will be the case when among-site rate variation is ignored (Fig. 1) (Waddell, 1995; Yang, 1996) and very likely explains the basal position of hedgehog in Cao and co-workers' (1997) analyses of the mtDNA data (those authors also assumed equal rates for all amino acid positions within each gene but allowed a different uniform rate for each gene). Even when heterogeneous rates are accommodated, the hedgehog can be placed on the tree in several locations without changing the likelihood significantly (Table I, Fig. 2A). The hedgehog therefore appears to represent a "rogue" taxon that cannot be placed reliably with these data and that possibly confounds attempts to estimate the relationships among the remaining taxa.

With the hedgehog/opossum long-branch attraction broken up (Fig. 2A), the unrooted tree for placental mammals (the ingroup topology obtained by pruning the opossum lineage from Fig. 2A) is consistent with rodent monophyly. The issue reduces to how the ingroup topology of placental mammals is rooted, that is, the reliability of the placement of the outgroup (the opossum sequence). In other molecular studies (e.g., Stanhope *et al.*, 1992), variable sites in opossum sequences have been shown to be essentially randomized relative to placental mammals, rendering the opossum sequences unlikely to provide a reliable root for those data. To examine this possibility specifically with respect to these data, we rooted the ingroup topology in Fig. 2A with 100 random sequences under the parsimony criterion (after exclusion of the hedgehog, to avoid the effect of long-branch attraction). These sequences were generated with MacClade (Maddison and Maddison, 1992) and were constrained to have the same base frequencies as the opossum sequence. Ninety-eight of those random outgroups rooted the tree within Rodentia, and many (24) rooted the tree at the same location as in D'Erchia and co-workers' (1996) analyses. The observation that random sequences nearly always root the tree at or near the same location as the opossum sequence raises the possibility of a spurious rooting in D'Erchia and co-workers' analyses. This tendency of the opossum



**Fig. 2.** (A) Maximum-likelihood tree ( $\ln L = -45,570.78$ ) for first and second codon positions for the "protein supergene" data of D'Erchia *et al.* (1996) estimated using a prerelease version of the PAUP\* 4.0 (GTR+I+ $\Gamma$  model of evolution,  $p_{inv} = 0.4103$ ,  $\alpha = 0.8239$ ). The hedgehog branch can be reconnected to each of the numbered branches with little change in likelihood score (see Table I). (B) The most likely tree under the constraint of rodent monophyly ( $\ln L = -45,579.04$ ). The root (attachment point of the opossum sequence) could be placed along the three numbered branches with insignificant change in likelihood (see Table I); attachment to the branch labeled one results in the identical topology as in A. The likelihood score of tree B does not differ significantly from that of the most likely tree (A).

**Table I.** Cost (Decrease in Log-Likelihood) Associated with Alternative Placements of Hedgehog on Tree A in Fig. 2 and Alternative Attachments of Opossum Outgroup on Tree B in Fig. 2<sup>a</sup>

Tree A			Tree B		
Branch	Cost	<i>P</i>	Branch	Cost	<i>P</i>
1	7.41	0.42	1	Best <sup>b</sup>	—
2	12.82	0.38	2	8.26	0.19
3	6.23	0.63	3	24.54	0.10
4	2.29	0.86			
5	5.93	0.56			
6	7.79	0.42			
7	6.40	0.40			

<sup>a</sup>Rearrangements that are not significantly different based on the Kishino–Hasegawa test.

<sup>b</sup>Attachment of the opossum sequence to branch 1 of Tree B produces Tree A.

sequence toward random rooting behavior, combined with the long-branch attraction between hedgehog and opossum sequences, contributed to their apparent strong refutation of rodent monophyly.

#### *Rodent Monophyly Is Not Significantly Refuted*

More importantly, with all taxa included and the analysis conducted under the more appropriate (GTR + I +  $\Gamma$ ) model, the likelihood score of the best rodent monophyly tree (Fig. 2B) is only 0.02% worse than that of the maximum-likelihood tree (Fig. 2A). This is not a significant difference as judged by the Kishino–Hasegawa test ( $P > 0.1$ ). Interestingly, when the Kishino–Hasegawa test is applied to the constrained (for rodent monophyly) and unconstrained trees under the (inappropriate) equal-rates GTR model (used by D’Erchia *et al.*, 1996), the best rodent monophyly tree is significantly worse than their tree ( $P = 0.012$ ). Thus, the apparent strong support for nonmonophyly of rodents reported by D’Erchia *et al.* (1996) is attributable to their use of an oversimplified model, that is, their incorrect assumption of equal rates across sites.

The results of the simulation analyses are similar. When the best rodent monophyly tree (Fig. 2B) is used as the model (true) tree and 100 data sets are simulated on that tree, maximum-likelihood analysis of the resulting data sets (under the model used to generate the data, GTR + I +  $\Gamma$ ) supports rodent monophyly only 75% of the time. That is, in 25% of the simulations, a nonmonophyletic Rodentia is supported, even though rodents are monophyletic on the model tree. We therefore cannot reject the null hypothesis ( $P = 0.25$ ) of no significant difference between the topologies presented in Fig. 2. Interestingly, when the simulated data sets are analyzed under parsimony, the probability of incorrectly inferring nonmonophyly of Rodentia increases to 0.68, and when analyzed using an equal rates likelihood model (GTR with equal rates), that probability increases to 0.85. In the oversimplified analyses, there is actually a higher probability of inferring the wrong tree than the model (true) tree. This analysis further demonstrates that the strong apparent support for rodent nonmonophyly reported by D’Erchia *et al.* (1996) resulted from ignoring rate heterogeneity across sites.

### The Utility of Models

Our analyses using the more appropriate heterogeneous-rates model (GTR+I+ $\Gamma$ ) not only break up the long-branch attraction between the hedgehog and the opossum sequences, but demonstrate that the statistical support (relatively high bootstrap values) that D'Erchia *et al.* (1996) reported for rodent nonmonophyly is an artifact of systematic error associated with ignoring among-site rate variation. This illustrates the point that large data sets (with respect to the number of bases) are not immune to systematic error. In fact, because inconsistent phylogenetic methods will (by definition) ascribe increasing confidence to incorrect estimates of topology as sequence length increases, the match between model and data becomes more critical for very long sequences rather than less so, as intuition might suggest. Thus, future analyses of complete mitochondrial genomes that fail to accommodate rate heterogeneity explicitly (e.g., Janke *et al.*, 1996; Cao *et al.*, 1997) will be susceptible to the same systematic error that misled D'Erchia *et al.* (1996).

Thus, these results demonstrate that rodent monophyly clearly is not refuted by the mtDNA genome data. In spite of the large number of base pairs, a more thorough sampling of taxa will be required to test adequately the hypothesis of rodent monophyly. In particular, more rodent and insectivoran sequences will be required (e.g., Nedbal *et al.*, 1996), and inclusion of xenarthran sequences (e.g., armadillo or sloth), the probable outgroup to the rest of the placental mammals (e.g., McKenna, 1975; Novacek, 1990), will divide the long branch leading to the opossum and thereby possibly provide a more reliable root for nonxenarthrous eutherians. Additional complete mtDNA sequences will soon be available. Analyses of these new data under appropriate models may or may not support D'Erchia and co-workers' (1996) conclusions; for the moment the molecular data do not support the dissolution of Rodentia, whose monophyly remains strongly supported by morphological analyses (Luckett and Hartenberger, 1993) and by at least some molecular studies (e.g., Porter *et al.*, 1996).

### APPENDIX

Table A1. Comparison of Models of Sequence Evolution Examined in this Paper"

	Base frequency	Substitution type	Rate heterogeneity	$\ln L$
JC	Equal	One	None	-50,999.91710
JC+I	Equal	One	Invariable sites	-47,911.44192
JC+ $\Gamma$	Equal	One	Gamma-distributed rates	-47,607.65530
JC+I+ $\Gamma$	Equal	One	Invariable sites + gamma-distributed rates	-47,605.36590
K2P	Equal	Two	None	-50,016.04069
K+I	Equal	Two	Invariable sites	-46,833.56109
K+ $\Gamma$	Equal	Two	Gamma-distributed rates	-46,471.74497
K+I+ $\Gamma$	Equal	Two	Invariable sites + gamma-distributed rates	-46,446.36463
HKY-85	Empirical	Two	None	-49,433.62471
HKY+I	Empirical	Two	Invariable sites	-46,201.59995
HKY+I+G	Empirical	Two	Gamma-distributed rates	-45,803.54994
HKY+I+ $\Gamma$	Empirical	Two	Invariable sites + gamma-distributed rates	-45,773.48647



Table AI. Continued

	Base frequency	Substitution type	Rate heterogeneity	ln L
GTR	Empirical	Six	None	-49,028.13801
GTR+I	Empirical	Six	Invariable sites	-45,965.65841
GTR+ $\Gamma$	Empirical	Six	Gamma-distributed rates	-45,620.12374
GTR+I+ $\Gamma$	Empirical	Six	Invariable sites + gamma-distributed rates	-45,593.28162

"All these models assume stationarity of each parameter. Likelihood scores are calculated on the tree of D'Erchia *et al.* (1996).

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